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# The Distributional Effects of the Hilmer Reforms on the Australian Gas Industry

by

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# THE DISTRIBUTIONAL EFFECTS OF THE HILMER REFORMS ON THE AUSTRALIAN GAS INDUSTRY

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### Abstract

We analyse changes in the Australian gas industry during 1990s that were driven by the Hilmer Reforms. We estimate the direct and indirect effects on household income of these gas industry changes by combining a computable general equilibrium model with a microsimulation model in a two-stage simulation procedure. The changes lead to minor effects on household income in all regions due to the unimportance of the gas industry at that time. Some regions benefit from the changes and some lose. Income inequality is only slightly affected by the changes.

JEL codes: C68, C69, L94, D31. Keywords: computable general equilibrium, gas, household income distribution, microeconomic reform, microsimulation.

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The framework applied here is an extension of that developed in Verikios and Zhang (2005, 2008) and initially applied in Verikios and Zhang (2010). The views expressed here are the authors' and do not necessarily reflect those of the Productivity Commission or Monash University. Thanks are due to Ken Clements and Ken Pearson for helpful comments on this work.

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### 1. Introduction

In the early 1990s Australian governments introduced a series of microeconomic reform policies for infrastructure industries (e.g., gas, ports, telecommunications, etc.); Productivity Commission (PC) (2002) summarises these reforms. The reforms were part of the process produced by the Hilmer Report and, subsequently, the National Competition Reform Act 1995 and the Competition Principles Agreement between Australian governments. The Hilmer Report's terms of reference focused on government businesses and regulations that had created protected enterprises: these had been a feature of industry policy in Australia for most of the 20<sup>th</sup> century. Hilmer argued for the introduction of competition policy in these areas in order to promote competition for the purpose of promoting community welfare, i.e., economic efficiency and other social goals (King and Maddock 1996). Thus, a major aim of the policy initiatives was to bring about market competition that, in turn, would lead to productivity improvements and attendant increases in real incomes, as well as better choice and services for consumers. Early in the reform process the Industry Commission (IC) estimated that the reforms could increase national output by around 5.5% of its current value at the time (IC 1995).<sup>1</sup> Since the initial introduction of the reforms, the affected industries have undergone significant structural changes that are observable in their cost structure and output prices. PC (2002) documents some of the infrastructure price changes in different Australian regions over the period 1990–91 to 2000–01.

As major service providers, changes in infrastructure industries can potentially have farreaching impacts on other industries, businesses and households. Both PC (1999) and Madden (2000) noted that the competition policy reforms were regarded by many in the community as being responsible for the increased economic divide between capital cities and regional Australia. Related to this, there has also been natural community concern over the impact on income distribution of sectoral changes, in general, and infrastructure industry changes, in particular, viewed as a result of the microeconomic reforms.

There is a paucity of Australian studies that have analysed the distributional effects of the Hilmer reforms with only two notable exceptions. In PC (1996a), an input-output model and household survey data are used to estimate the effects on household expenditure of price reforms by government trading enterprises (GTEs) in the electricity industry and the water, sewerage and

<sup>&</sup>lt;sup>1</sup> Filmer and Dao (1994) estimated that a wider package of microeconomic reforms, including those related to infrastructure industries, could raise GDP by between 12% and 15%.

drainage services industry. In a companion paper (PC 1996b), a more sophisticated approach is adopted. A computable general equilibrium (CGE) model in conjunction with an income distribution model is used to analyse the effects of a specific set of reforms on the sources of household income. Each of these studies concentrates on only one side of the household budget, so the overall impact on household real income remains unclear. Moreover, input-output models, as applied in PC (1996a), do not capture effects generated from sectoral reallocation of resources that are arguably the most important effects of any policy change. As a result, the effects of a policy change derived from such a model may be misleading.

As a response to the shortcomings of previous studies we conduct a more comprehensive analysis of the effects of industry changes on household income distribution. We analyse the distributional effects of sectoral changes that have occurred at the same time as the implementation of microeconomic reform policies, by integrating both the income and expenditure sides of the household budget to capture the total (direct and indirect) effect on household real income. An ideal approach to such an analysis is to use a CGE model directly incorporating individual households. Unfortunately, such a highly disaggregated multi-household model is unavailable for Australia.<sup>2</sup> Data limitations are a serious impediment to the development of such a model.

Given these limitations, we adopt a simple but informative two-step approach in which a multi-region CGE model is first simulated by gas-industry-specific changes during 1990s to generate aggregate changes in the prices of goods and services, and productive factor returns. The resulting aggregate effects are then applied to a separate and highly disaggregated microsimulation model for a detailed analysis of changes in individual household expenditure and income. Thus, this work contributes to the limited analysis of the distributional effects of the microeconomic reforms motivated by the Hilmer Report.

<sup>&</sup>lt;sup>2</sup> Regardless, there has been some progress in recent years in this area for other countries, e.g., for the US (Slemrod 1985), for Madagascar (Cogneau and Robilliard 2000), for the UK (Plumb 2001), for Nepal (Cockburn 2006), and for the Philippines (Cororaton and Cockburn 2007).

### 2. Microeconomic reform during the 1990s

### 2.1 Australian infrastructure industries and the Hilmer Reforms<sup>3</sup>

At the beginning of the 1990s Australian governments began an extensive process of microeconomic reform of Australian infrastructure industries, including electricity, gas, water, sewerage, urban passenger transport, port services, rail freight, telecommunications and postal services. The main objectives of these reforms were to increase competition and performance in these industries, and thus bring about higher living standards.

Prior to the commencement of the reform process almost all infrastructure industries were dominated by GTEs providing services with monopoly rights. Thus the reform process has been largely concerned with improving the performance of GTEs. With respect to GTEs, the reform process can be categorised into four broad areas: commercialisation; corporatisation; capital market disciplines; and competition policy.

*Commercialisation*. This involves GTEs taking a more market-driven approach to service provision and pricing. To aid the commercialisation process, competitive tendering and contracting out of service provision have been introduced, community service obligations are now funded in a more direct and transparent way, and GTE regulatory functions have been transferred from GTEs to independent regulators.

*Corporatisation*. This focuses on making GTEs autonomous entities, within the public sector, with commercially-oriented boards pursuing commercial objectives without ministerial interference. Financial and non-financial performance monitoring and reporting regimes were set up to measure and compare performance. Price regulation has also been largely transferred from ministerial control to independent regulators.

*Capital market disciplines*. Traditionally, GTEs were not required to earn a commercial rate of return on their assets in the way that private sector firms must. This has now changed, with many governments requiring GTEs to either reduce negative rates of return or earn higher positive rates of return.

*Competition policy*. The implementation of the National Competition Policy Agreement has focussed on removing existing entry barriers to infrastructure industries and thereby stimulating competition and increasing contestability. Increased competitive pressure is aimed at lowering prices and increasing service provision and quality.

 $<sup>^{3}</sup>$  This section draws on PC (2002), section 1.3.

### 2.2 The Australian gas industry and the Hilmer Reforms<sup>4</sup>

For historical and cost reasons, not all Australian households and industries use natural gas; at the end of last century only 47 per cent of Australian households were connected to natural gas. This figure varies widely across regions. The historical development of transmission and distribution networks in each region determines the degree of access by households and industries. Further, less densely populated areas face higher distribution costs, which also limits the degree of access to gas.

For most of the twentieth century Australian governments owned gas utilities that controlled gas pipelines, distribution networks and retail businesses. In 1994, the Council of Australian Governments made a commitment to "free and fair trade in natural gas", which marked the beginning of gas industry reforms. Following this, publicly owned gas utilities were either corporatised or privatised. Policy and regulatory impediments to interstate trade and retail competition were gradually reduced or removed.

As publicly owned gas utilities were corporatised or privatised, widespread structural changes took place in the gas industry across regions. Vertically integrated gas transmission and distribution activities were separated. Employment levels fell and work practices changed. The reduction in employment was accompanied by contracting out of services by gas utilities.

To introduce competition in gas trade, gas prices were gradually deregulated. More customers were given the freedom to choose their gas retailer. In some regions, price rebalancing between customer classes was also undertaken to make gas prices more reflective of the costs of supplying different customer types and the price sensitivity of their demand.

### 3. Method

Our modelling approach was developed and initially applied in Verikios and Zhang (2010). It links two separate analytical frameworks for the purpose of generating results at a high level of household detail without a complex CGE model that fully integrates individual households.

### 3.1 The history of linked models

As the inventor of microsimulation, it is not surprising that Orcutt (1967) was the first to describe a process for linking models that operate at differing levels of aggregation. He

<sup>&</sup>lt;sup>4</sup> This section draws on PC (2002), chapter 3.

envisaged multiple models, each describing part of the economy, being linked as modules that together would describe the overall system. The most succinct summary of alternative approaches to linking micro and macro models is provided by Bækgaard (1995) who identifies the following methods:

- 1. a top-down approach in which the micro model is adjusted to match an exogenous macro aggregate;
- 2. a bottom-up approach in which a change generated in the micro model is used to adjust the macro model;
- 3. a recursive linkage approach in which there is a two-way lagged interaction between models; and
- 4. an iterative approach in which the two models are solved simultaneously within each period.

A fifth approach proposed by Toder et al. (2000) involves the micro and macro models being solved separately over the full simulation period, with the models then calibrated and resolved until convergence is achieved. A further alternative is to build a model that inherently includes both a micro and macro dimension (Davies 2004). In principle, such a fully integrated model is preferred; in practice, most models in the literature take a recursive-linkage approach.<sup>5</sup> This reflects the practical difficulties of including both dimensions within the one model.

A common feature of linked CGE-microsimulation models developed to date is a focus on labour supply (e.g., Aaberge et al. (2007), Arntz et al. (2008), Fredriksen et al. (2007). Recent examples of linked models developed to assess distributional issues include Herault (2006) and Herault (2007). Within Australia, there have been only limited attempts to link micro and macro models. As far as we are aware, the earliest Australian example is provided by Meagher and Agrawal (1986) in which output from a CGE model was used to reweight the 1981–82 National Income and Housing Survey. Their approach was updated by Dixon et al. (1996), who also foreshadowed an iterative linking of a CGE model to either a static or dynamic microsimulation model. In related work, Polette and Robinson (1997) used the top-down approach to link an aggregated version of the MONASH dynamic CGE model to a microsimulation model of the Australian income support system.

<sup>&</sup>lt;sup>5</sup> See footnote 2 for some examples of fully integrated models.

Of the two Australian studies that have analysed the distributional effects of the Hilmer reforms, PC (1996b) follows the pioneering work (in the Australian context) of Meagher and Agrawal (1986) by using a CGE model in conjunction with an income distribution model to analyse the effects of some of the Hilmer reforms on the sources of household income. PC (1996a) applies an input-output model and household survey data to estimate the effects on household expenditure of price reforms by GTEs in the electricity industry and the water, sewerage and drainage services industry. But input-output model are inappropriate for analysing distributional effects.<sup>6</sup> Each of these studies concentrates on only one side of the household budget, so the overall impact on household real income is unclear.

### 3.2 Analytical framework: a linked CGE-microsimulation top-down approach

Most of the Australian studies mentioned above have focused on linking a CGE model to a detailed microsimulation model of household income. Thus, they have mostly ignored the differences in expenditure patterns across households and their effect on estimates of distributional effects. As a response to these shortcomings, we develop a more comprehensive framework for analysing distributional issues by integrating both sides of the household budget to capture the direct and indirect effects on household real income.<sup>7</sup> We do this by adopting the top-down approach. That is, a multi-region CGE model – the Monash Multi-Region Forecasting (MMRF) model (Naqvi and Peter 1996) – is first simulated using industry-specific changes to generate aggregate changes in the prices of goods and services, and factor returns. The resulting aggregate effects are then applied to a separate and highly disaggregated microsimulation model – the MMRF Income Distribution (MMRF-ID) model – for a detailed analysis of changes in individual household expenditure and income; Figure 1 gives a diagrammatic representation of the analytical framework.

<sup>&</sup>lt;sup>6</sup> Input-output models assume all prices, including factor prices, are fixed. In reality, any reallocation in resources across sectors will alter factor prices and incomes. Thus, using input-output models to analyse structural change rules out, by assumption, any effects on factor prices and income.

<sup>&</sup>lt;sup>7</sup> A change in a given infrastructure industry fundamentally affects household real income in two ways. First, a change in the price of the service will affect household real expenditure. This is usually referred to as the 'direct effect'. Second, a price change could be the result of changes in a service provider's cost structure, such as raw material usage, primary factor inputs or new technologies. Such changes alter factor returns and the income of factor owners, i.e., households. Further, the prices of other goods and services may also be affected. These effects are usually referred to as 'indirect effects'.

### Figure 1 Linkages between the MMRF and MMRF-ID models



 $\mathbf{a}_{\mathrm{Monash\,Multi-Region\,Forecasting\,model.}} \mathbf{b}_{\mathrm{MMRF\,Income\,Distribution\,model.}}$ 

### 3.3 The MMRF model

### 3.3.1 A linear equation system

The MMRF model is represented by equations specifying behavioural and definitional relationships. There are m such relationships involving a total of p variables and these can be compactly written in matrix form as

$$A\boldsymbol{\nu} = \boldsymbol{0}, \tag{1}$$

where A is an  $m \times p$  matrix of coefficients, v is a  $p \times 1$  vector of *percentage* changes in model variables and **0** is the  $p \times 1$  null vector. Of the p variables, e are exogenous (e.g., taxes). The e variables can be used to shock the model to simulate changes in the (p-e) endogenous

variables. Many of the functions underlying (1) are highly nonlinear. Writing the equation system like (1) allows us to avoid finding the explicit forms for the nonlinear functions and we can therefore write percentage changes (or changes) in the (p-e) variables as linear functions of the percentage changes (or changes) in the *e* variables. To do this, we rearrange (1) as

$$A_n \boldsymbol{n} + A_x \boldsymbol{x} = \boldsymbol{0}, \tag{2}$$

where *n* and *x* are vectors of percentage changes in endogenous and exogenous variables.  $A_n$  and  $A_x$  are matrices formed by selecting columns of *A* corresponding to *n* and *x*. If  $A_n$  is square and nonsingular, we can compute percentage changes in the endogenous variables as

$$\boldsymbol{n} = -\boldsymbol{A}_n^{-1}\boldsymbol{A}_x \boldsymbol{x}. \tag{3}$$

Computing solutions to an economic model using (3) and assuming the coefficients of the *A* matrices are constant is the method pioneered by Johansen (1960).

Equations (1) represent the percentage-change forms of the nonlinear functions underlying the model; these forms are derived by total differentiation. Thus, (1) is an approximation based on marginal changes in the independent variables. So (3) only provides an approximate solution to the endogenous variables n; for marginal changes in x the approximation is accurate but for discrete changes in x the approximation will be inaccurate. The problem is the standard one of numerical integration.

The problem of accurately calculating n for large changes in x is equivalent to allowing the coefficients of the A matrices to be nonconstant. The problem is solved by breaking the change in x into i changes. The multistep solution procedure requires that there are (i-1) intermediate values of the underlying (levels) values of n, i.e., N. The intermediate values of N are obtained by successively updating the values of N after each of the i steps is applied. Once the values of N are updated for any given step, the coefficients of the A matrices in (3) are recomputed before (3) is solved again.<sup>8</sup>

### 3.3.2 Theory

The MMRF model describes the supply and demand side of commodity and factor markets in the eight Australian states and territories. Each region contains five types of representative

<sup>&</sup>lt;sup>8</sup> The model is implemented and solved using the multistep algorithms available in the GEMPACK economic modelling software (Harrison and Pearson 1996).

agent – producers, physical capital investors, households, governments and foreigners. In the version applied here, there are fifty-four producers or industrial sectors in each region, each producing one commodity. Commodities are traded between regions and are also exported. There is a single representative household, and nine government sectors (eight regional and one national). Foreigners supply imports to each region at fixed c.i.f. prices, and demand commodities (exports) from each region at variable f.o.b. prices.<sup>9</sup>

Regional supplies of and demands for commodities are derived from optimising behaviour of agents operating in competitive markets. Producers employ constant-returns-to-scale technology and combine primary factors (land, labour and capital) and intermediate inputs using a series of nested Leontief and CES production functions; see Figure 2. Land use is confined to, and fixed within, the agricultural sectors.

In this comparative static version of MMRF the national supply of capital is fixed whereas in reality capital can vary in the long run. But this study is concerned with the reallocation of existing factors rather than growth effects. This means that any excess demands for capital at initial prices (due to gas industry changes) are partly reflected in rental price changes and partly reflected in the reallocation of capital across regions and sectors: capital moves between industries and across regions to maximise its rate of return.

MMRF specifies the labour input into the primary factor bundle as a constant-elasticity-ofsubstitution (CES) combination of eight occupational categories.<sup>10</sup> In original form, the CES prices for these occupational categories are identical. Thus, the CES demand function for labour operates like a Leontief (fixed proportions) production function. The MMRF-ID model (see Section 3.4) specifies labour income for households according to these occupational categories. In order to exploit the richness of the labour income data in MMRF-ID, the MMRF model is modified to allow for an occupation-specific price of labour in each region. This allows MMRF to use the occupational wage bill data for each industry to differentially affect demand for, and

<sup>&</sup>lt;sup>9</sup> Thus, MMRF assumes export commodities are imperfectly substitutable with exports from other countries in the tradition of Armington (1969). This formulation means the terms-of-trade are endogenous and Australian regions are modelled as 'almost small'.

<sup>&</sup>lt;sup>10</sup> The occupational grouping used is the Australian Standard Classification of Occupations (first edition) (ABS 1986).





Source: Peter et al. (1996).

therefore the price of, each occupational category. To implement occupation-specific wage rates, the MMRF theory and data relating to the labour market is modified [see Verikios and Zhang (2005) for further details]; below we describe further changes made for this work.

On the demand side of the labour market, the price paid by firms in industry *j* for labour of (occupation) type *m* in region *r*,  $PL_{imr}$ ,

$$PL_{jmr} = W_{jmr}T_{jr}, (4)$$

where  $W_{jmr}$  is the wage received by workers and  $T_{jr}$  is the power of the payroll tax rate paid by firms in industry *j* in region *r*, i.e.,  $T_{jr} = 1 + t_{jr}$  where *t* is the tax rate. In the original model,  $W_{jmr}$  is only defined over industries and regions. For a given  $PL_{jmr}$ , firms will determine their cost-minimising demand for each labour type.

On the supply side, the supply of labour type *m* in region *r*,  $LS_{mr}$ ,

$$LS_{mr} = RW_{mr}^{\beta}, \tag{5}$$

and

$$RW_{mr} = \frac{W_{mr}}{CPI_r},\tag{6}$$

where  $W_{mr}$  is the average wage paid to labour type *m* in region *r*, and  $CPI_r$  is the consumer price index in region r. Thus, the supply of each labour type is a positive function of the real wage,  $RW_{mr}$ , and  $\beta$ , the labour supply elasticity.  $\beta$  is set at 0.15 reflecting econometric evidence on labour supply in Australia (Kalb 1997). In the original model,  $LS_{mr}$  is only defined over regions.

Labour market equilibrium (including unemployment) is determined by imposing a relation between real wages and employment  $L_{mr}$  of the form,

$$L_{mr} = RW_{mr}^{1/\gamma}, \tag{7}$$

where  $1/\gamma$  represents the real wage elasticity of employment, i.e., the responsiveness of employment to changes in the real wage. In any perturbation of the model,  $1/\gamma$  determines the degree to which increases (decreases) in the demand for the *m*-, *r*-th labour type will be reflected in higher (lower) employment or in a higher (lower) real wage. We calibrate  $\gamma$  so that employment responses are half as large as real wage responses. This value is based on casual empiricism of the Australian labour market where real wages grow faster than employment; regardless, we conduct sensitivity analysis with respect to the value of  $\gamma$ .

The national consumer price index (CPI) is the numeraire, thus nominal price changes are measured relative to this composite price.

### 3.4 The MMRF-ID model

MMRF-ID is a microsimulation model that represents the distribution of real incomes across households in the eight Australian regions; it only comprises equations specifying definitional relationships.

### 3.4.1 Theory

Two measures commonly used to compute the benefits that accrue from a price change are compensating variation (CV) and equivalent variation (EV). Both compute the amount of money that would bring the consumer back to their original utility level prior to a price change. The CV values this amount at new prices while the EV values it using original prices. Consumer surplus is a related measure. Both CV and EV apply a 'money-utility' concept rather than utility itself.

A modified version of the CV is based on redefining real income as constant purchasing power; i.e., the amount of money that allows the consumer to purchase the same bundle of goods as before the price change. Applying this concept to measure changes in real income means there is no need to make any specific assumptions about consumer preferences or utility functions. This approach is in line with Slutsky's decomposition of price change effects.

The computation of CV normally assumes unchanging household income and, therefore, emphasises only the role of each household's different consumption patterns in determining the welfare impact of a price change. But in a general equilibrium framework household income is not constant. As a result, the above definition of CV can be usefully extended to account for changing income. Similar to the expenditure side, the income side of the modified CV can be interpreted as the amount of money that would encourage the household to supply the same amount of factors as prior to any price change.

For a household, real income can then be defined as nominal factor earnings and transfers received from different sources deflated by a household-specific consumer price index (HCPI). Then, the first-order approximation to the percentage change in the *i*-th household's CV, relative to the initial consumption bundle and factor ownership, can be expressed as

$$cv^{i} = -\left(w^{i} - p^{i}\right),\tag{8}$$

where  $w^i$  is the percentage change in income received by household *i*, and  $p^i$  is the percentage change in the HCPI for household *i*.  $p^i$  is the average of the percentage changes in the prices of goods and services, weighted by their expenditure shares,

$$p^{i} = \sum_{n} \theta_{n}^{i} p_{n} , \qquad (9)$$

where  $\theta_n^i$  is the *i*-th household's expenditure share for good *n*, and  $p_n$  is the percentage change in the price of good *n*.

Differences in the sources of income  $w^i$  for the *i*-th household can be expressed as

$$w^i = \sum_f \phi^i_f w_f , \qquad (10)$$

where  $\phi_f^i$  is the share of income source *f* in the *i*-th household's income, and  $w_f$  is the percentage change in the factor price of income source *f*. If the HCPI rises relative to income, compensation will be required  $(cv^i > 0)$  in order to keep the household at its initial level of consumption.

The income side of our modified CV is the amount of money that would encourage households to supply the same amount of factors as prior to any price change. But we know that the general equilibrium effects of industry changes will lead to changes in factor supply and employment as well as factor returns. To account for such changes, we redefine  $w^i$  as

$$w^i = \sum_f \phi^i_f w_f q_f , \qquad (11)$$

where  $q_f$  is the percentage change in the demand (or employment) of income source f. Thus, our modified CV assesses the impact of a policy change on a given household or household group via the computation of the change in real income.

In computing real household income changes in MMRF-ID,  $p_n$ ,  $w_f$  and  $q_f$  are set equal to the values generated in MMRF;  $\theta_n^i$  and  $\phi_f^i$  are calculated from the MMRF-ID database.

### 3.4.2 Data

The MMRF-ID data are based on unit-record household survey data taken from the 1993– 94 Household Expenditure Survey (HES93) (ABS 1994). The survey contains detailed information on household consumption patterns and income sources of 8,389 sample households in existence around the beginning of the 1990s across the eight Australian states and territories. On the income side, the HES93 lists not only private income sources, such as wages and salaries from eight occupations and non-wage income from investment or business sources, but also various government transfer payments, such as family allowances, unemployment benefits and age pensions (see Table 1). It also contains detailed expenditure data on more than 700 goods and services. These items are aggregated to 54 groups consistent with the commodities in MMRF. This information is used to calibrate the MMRF-ID database.

MMRF model	MMRF-ID model								
Labour income sources	Managers, Professional, Para-Professional, Trades Persons, Clerks, Sales Persons Plant/Machine Operators, Labourers.								
Non-labour income sources	Interest, Investment, Property Rent, Superannuation, Business, Workers Compensation, Accident Compensation, Maintenance, Other Regular Sources, Private Scholarship, Government Scholarship, Overseas Pensions.								
Government benefits	Sickness Benefits, Family Allowance, Veteran's Pensions, Unemployment Benefits, Age Pensions, Widows Pensions, Disable Pensions, Sup Par Benefits, Wife's Pensions, Other Australian Government Benefits, AUSTUDY Support, Carer's Pensions, Other Overseas Government Benefits.								
Income tax	Direct tax.								

Table 1 Mapping between household income sources in MMRF and MMRF-ID

In reporting distributional effects from MMRF-ID, we group households according to regional income deciles. Given the focus of this work is the effect of gas industry changes, Table 2 presents the national share of household expenditure allocated to gas across income deciles. As expected, the share falls as household income rises reflecting the very stable nature of demand for gas with respect to income, i.e., a low income elasticity of demand. Nevertheless, the expenditure shares vary significantly across region with the smallest shares in Tasmania and the highest shares in Victoria. Table 2 also presents the distribution of household income across income sources for each decile. It shows that government benefits are the dominant source of household income for the first three deciles, whereas labour income is the most important income source for the remaining seven deciles. The data also show a steadily rising direct tax rate as income rises. The data patterns are as expected.

Income	Share of gas	Sh	are of household	income by incom	ne source	
decile	expenditure in total	Non-labour	Labour income	Government	Direct	Total
	expenditure	income		benefits	taxes	
Lowest	0.008	-0.125	0.318	0.778	0.029	1.000
Second	0.007	0.074	0.393	0.476	0.057	1.000
Third	0.008	0.124	0.318	0.498	0.060	1.000
Fourth	0.007	0.084	0.459	0.373	0.085	1.000
Fifth	0.006	0.143	0.556	0.179	0.122	1.000
Sixth	0.006	0.119	0.633	0.106	0.142	1.000
Seventh	0.006	0.128	0.651	0.066	0.156	1.000
Eighth	0.005	0.105	0.700	0.026	0.168	1.000
Ninth	0.005	0.098	0.704	0.014	0.184	1.000
Highest	0.004	0.147	0.621	0.006	0.225	1.000

 Table 2 Gas expenditure and income source shares in MMRF-ID, national (fraction)

Source: MMRF-ID database.

### 4. Calculating gas-industry-specific changes

Determining changes specific to the gas industry over the 1990s is an important input to this work. The structure of the gas industry at the end of the 1990s was different from that at the beginning of the microeconomic reform process in the early 1990s. While it seems reasonable to attribute most of these changes to the reform process directed specifically to the gas industry, changes have occurred in the other parts of the economy that are likely to also have influenced the changes observed in the gas industry.<sup>11</sup> But we do not wish to consider all historical events that have reshaped the gas industry over the 1990s, but rather to isolate all gas-industry-specific changes. To estimate such changes, the observed changes in the gas industry need to be adjusted to remove the effects of external factors. If complete information on changes in the quantities of industry inputs and outputs was available, these changes could be imposed directly as shocks in the model to generate the requisite equilibrium prices and quantities for gas, as well as other commodities and primary factors. But information is only available on two industry variables: employment and output prices.

The observed changes in industry gross employment contain an expansionary effect caused by economy-wide output growth (due to changes in productivity, tastes and preferences, technology, etc.), which may be unrelated to industry-specific changes. To remove this effect, employment per unit of output is used to simulate the change in the gas industry's employment.

<sup>&</sup>lt;sup>11</sup> See Quiggin (1997) for an *ex ante* analysis of the effects of microeconomic reform on the gas industry.

Employment per unit of output is calculated as observed gross employment divided by the quantity of output, or more precisely, employment per petajoules supplied.

In imposing the changes in employment per unit of output on the CGE model, this typically endogenous variable must bet set as exogenous. This is accommodated by setting labouraugmenting technical change as endogenous. This implicitly assumes that any change in unitoutput employment can be attributed to a change in industry-specific labour productivity.

In calculating the gas price shocks, we want to remove the effects of non-gas-industry factors, e.g., inflation, income growth, population expansion, etc. The impacts of these external effects on the price of gas can be removed, to a large extent, by calculating a 'real price index', i.e., the observed market price divided by the consumer price index (CPI). If the CPI is taken as a proxy for the price index of all goods and services, the real price of gas can be conveniently interpreted as a relative price. Any deviation of the real price from the CPI can then be interpreted as indicating changes caused purely by gas-industry-specific factors.

The real price of gas is typically an endogenous variable in a CGE model. To impose the price change in MMRF, we set it as exogenous and all-input-augmenting technical change is set as endogenous. This implies that any price change can be attributed to a change in the technology affecting the use of all inputs in the production of gas.

Like many infrastructure industries, the gas industry charges different prices for different customers. For instance, at least three sets of prices are reported by gas firms: residential, commercial and industrial (see ABS 2001a). Over the 1990s, the basic cost structure of supplying and distributing gas changed, along with prices for different customers. Price data indicates that gas firms have rebalanced prices for different customer groups (PC 2002). To account for price rebalancing the MMRF model is modified as follows. In each region, instead of shocking a single price of gas, two price shocks are introduced: one for producers who use gas as an intermediate input, and the other for households who use it for consumption. The supply price of gas by the industry is then a weighted average of the two prices, which is set equal to the total cost of production.<sup>12</sup>

Changes in the gas industry are also likely to affect government revenue. To neutralise the effect of changes in government revenue in the analysis, we fix the federal budget deficit and endogenise the income tax rate. We also fix the budget deficit for all state governments and

<sup>&</sup>lt;sup>12</sup> Zero pure profits in production is assumed for all industries.

endogenise their payroll tax rates. This assumes that for a given level of public expenditure, any increased (decreased) tax revenue due to the changes in the gas industry will be automatically recycled to households though a decrease (increase) in their income tax rates, and higher (lower) pre-tax wage rates due to lower payroll tax rates on firms. These assumptions are likely to be important in determining how any efficiency gains are distributed across households, so we conduct sensitivity analysis where these assumptions are varied. To explore the effects of parameter uncertainty, we also conduct sensitivity analysis with respect to a number of important model parameters.

On the government expenditure side, real government consumption expenditure is a fixed share of real household consumption expenditure; this is also subject to sensitivity analysis where we fix real consumption by all governments. In turn, household consumption expenditure is fixed share of household disposable income. Similarly, government investment expenditure is a fixed share of total (private and public) investment expenditure. Private investment expenditure moves in line with any changes in each industry's capital stock.

### 5. Results

### 5.1 Economy-wide effects

In this section the estimated changes in the real price and employment per unit of output in the gas industry are used to shock MMRF to project the aggregate effects of these changes on the general economy. The shocks are estimated from published statistics and are reported in Table 3.<sup>13</sup> We see that employment per unit of output decreased significantly in all regions for the period 1989–90 to 1999–00, from a maximum of -89% in Victoria to a minimum of -39% in the Northern Territory (NT). Real business prices fell in most regions, but only slightly, and real household prices showed no distinguishable pattern across regions, but they did increase significantly in South Australia (11%), Northern Territory (20%) and the Australian Capital Territory (15%).

<sup>&</sup>lt;sup>13</sup> See the notes to Table 3 for the sources used. Further details on the calculation of the shocks are presented in Verikios and Zhang (2005).

Variable	NSW	Vic	Qld	SA	WA	Tas <sup>b</sup>	NT <sup>c</sup>	ACT
Employment per unit of output	-76.7	-88.7	-86.3	-44.5	-42.7	na	-39.4	-93.1
Business prices <sup>a</sup>	-13.5	-1.7	-1.2	0.7	-5.7	na	1.7	-4.7
Household prices	2.1	-3.0	-9.5	11.2	-11.1	na	20.1	15.0

Table 3 Estimated changes in gas industry variables: 1989–90 to 1999–00 (percentage change)

Source: ABS (2001a, 2001b) and AGA (1991, 1992, 1994, 1995, 1996, 1997, 1998, 1999, 2000).

<sup>a</sup> Business prices were calculated as the consumption-weighted average of reported commercial and industrial prices in ABS (2001a). <sup>b</sup> Tasmania does not have a natural gas industry; gas consumption in Tasmania is in the form of liquefied petroleum gas (LPG). <sup>c</sup> The nominal gas prices for the NT reported in ABS (2001c) are calculated as the weighted average of the LPG price and the natural gas price. This is because the NT has only a small natural gas industry, and a large proportion of gas consumption in the NT is LPG. Thus the LPG price dominates the changes in the average gas price computed for the NT; LPG prices are heavily influenced by world petroleum prices. Ideally, only the change in the natural gas price would be applied in MMRF but this is not available.

A CGE model captures both the direct and indirect effects of a given shock to the economy. The major determinant of the direct effects of changes in the gas industry is its importance in the economy as a whole. Our model data indicates that gas output comprised around 0.3% of national output in 1993-94 (our base year); the share of gas output in total output was largest in Victoria (0.5%) and smallest in the ACT (0.03%). This suggests that changes in the gas industry will lead to small direct effects, but of varying magnitudes across regions.

The major determinant of the indirect effects of changes in the gas industry is its importance to other industries, as indicated by its sales as an intermediate input to production. Our model data indicates that intermediate input usage makes up around two-thirds of total gas sales, with household consumption comprising the remainder. This suggests that the indirect effects from changes in the gas industry will be stronger than the direct effects.

The results of applying the estimated changes in employment and prices to MMRF are reported in Table 4. The estimated changes in unit-output employment will determine the changes in labour productivity.<sup>14</sup> The estimated changes in business and household prices are aggregated in MMRF, weighted by their shares in total sales, to determine the changes in the basic price index of gas. In turn, the change in the basic price determines the change in the productivity of all inputs, i.e., all primary factors and intermediate inputs. The change in labour and all inputs productivity is summed to give an average productivity change. This change is closely related to the change in the basic price for the industry. Average productivity is projected

<sup>&</sup>lt;sup>14</sup> When referring to productivity changes in discussing model results, we are referring to the model equivalent of input per unit of output. Thus, a negative change in productivity represents an improvement.

to improve in New South Wales (NSW), Victoria, Queensland and Western Australia (WA); in these regions either both business and household prices fell in real terms, or their share-weighted average was lower. Average productivity is projected to deteriorate in South Australia (SA), the Northern Territory (NT), and the Australian Capital Territory (ACT); in these regions either both business and household prices rose in real terms, or their share-weighted average was higher. Note that the small size of the average productivity changes for the gas industry and the small size of the industry in overall activity suggests that the economy-wide effects are likely to be small.

 Table 4 Gas industry effects due to changes in unit-output employment and relative output prices between 1989–90 and 1999–00 (percentage change)

Variable	NSW	Vic	Qld	SA	WA	Tas	NT	ACT
Labour productivity <sup>a</sup>	-92.6	-98.6	-98.1	-68.7	-60.0	0.0	-64.9	-99.5
All inputs productivity <sup>a, b</sup>	8.7	22.0	12.2	21.9	1.1	0.0	24.0	42.5
Basic price	-9.6	-2.2	-3.6	4.6	-7.1	0.0	7.4	4.7
Average productivity <sup>a</sup>	-9.6	-2.2	-3.6	4.6	-7.2	0.0	7.3	4.7

Source: MMRF simulation.

a This is the input requirement per unit of output; thus, a negative sign signifies an improvement. b This relates to all primary factors and intermediate inputs.

The national changes in relative occupational incomes (Table 5) indicate which occupations are favoured by the gas industry changes; these show large relative reductions for Tradespersons, Clerks, and Labourers and related workers (LRW). This is because two-thirds of wage payments in the gas industry are made to these three occupations. Thus, when significant labour shedding occurs in this industry it is primarily Tradespersons, Clerks, and LRWs who are affected, and consequently the wage rates for these occupations must fall for them to be reemployed in other industries. Occupations that are least used in the gas industry experience the largest increases in relative incomes; these are Managers and administrators, and Salespersons and personal service workers.

Variable	NSW	Vic	Qld	SA	WA	Tas	NT	ACT	Aust
Labour income	0.07	-0.10	-0.06	-0.12	0.06	-0.03	0.10	-0.17	-0.02
Managers & administrators	0.22	0.27	0.06	0.09	0.07	0.01	0.09	-0.05	0.17
Professionals	0.11	-0.04	-0.06	-0.13	0.09	-0.03	0.13	-0.09	0.02
Para-professionals	0.08	-0.09	-0.06	-0.14	0.08	-0.01	0.13	-0.15	-0.01
Tradespersons	-0.04	-0.36	-0.14	-0.22	0.02	-0.07	0.07	-0.41	-0.15
Clerks	0.01	-0.20	-0.11	-0.17	0.04	-0.03	0.10	-0.15	-0.08
Salespersons & personal service workers	0.20	0.24	0.02	0.04	0.08	-0.02	0.07	-0.02	0.15
Plant & machine operators; drivers	0.09	-0.13	-0.05	-0.11	0.08	-0.03	0.14	-0.35	-0.01
Labourers & related workers	-0.07	-0.45	-0.13	-0.23	0.01	-0.02	0.06	-0.38	-0.18
Non-labour income	0.10	0.16	0.05	0.05	0.13	0.05	0.22	-0.13	0.10
Unemployment benefits	-0.14	-0.03	0.06	0.24	-0.11	0.07	-0.11	0.42	-0.03
Other government benefits	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0	0.0
Direct tax rate									0.01
Real household disposable income	0.03	-0.01	-0.02	-0.11	0.04	-0.03	0.02	-0.15	0.00
СРІ	0.03	-0.05	-0.02	0.04	0.02	0.02	0.06	0.02	0.00

Table 5 Regional effects of changes in the gas industry between 1989-90 and 1999-00(percentage change)

Source: MMRF simulation.

The national pattern of relative changes in occupational incomes is repeated at the regional level but with different absolute changes across regions. For instance, labour incomes rise in NSW, WA and NT, and fall in other regions. In general, the relative movements in labour income across regions reflect the relative productivity changes across regions; relative productivity improvements lead to higher relative labour incomes and vice versa. An exception is the NT, where labour income rises by more than in other regions. This reflects the general expansion in NSW and its higher demand for mining goods imported from NT.

Non-labour income also increases nationally reflecting increased demand for capital and land.<sup>15</sup> The relative changes in non-labour income across regions reflect the pattern of movements in labour income across regions. Unemployment benefits fall in regions that experience higher employment and rise in regions that experience lower employment.

Besides the changes in primary factor incomes, the direct tax rate will also affect household post-tax income. With the assumption of a fixed federal budget deficit and an endogenous direct tax rate, changes in the direct tax rate are driven by the effect of changes in the gas industry on total tax revenue. Changes in total tax revenue are driven by the effect of changes in the gas industry on the level of economic activity. While productivity improves in some regions and

<sup>&</sup>lt;sup>15</sup> Non-labour income is the sum of capital and land rentals. It is used in the MMRF-ID model to shock all nonlabour sources of household factor income.

falls in others, there is a small contractionary effect in net terms on economic activity nationally. Thus, tax revenue decreases and this is retained via a slightly higher direct tax rate (0.01%).

The changes in the gas industry affect not only different household income sources but also the prices of goods (and services). Given a household's preferred bundle of goods, changes in the prices of goods directly affect a household's expenditure. The regional CPI reported in Table 5 is the expenditure-weighted average of the prices of goods consumed in each region. The CPI effects indicate that the cost of household expenditure rises in all regions except Victoria and Queensland. The CPI effects reflect the average change in the basic price of goods in each region. But the changes in the gas industry have relatively little impact on other industries, so the changes in the basic price of goods largely reflect the change in the household price of gas in each region.

### 5.2 Household effects

The changes in the prices goods and factor incomes projected by MMRF are used as shocks to the MMRF-ID model to compute changes in individual household real income. These results are presented by income deciles for each of the eight regions and nationally in Table 6. Given that the economy-wide effects of the gas industry are small, we should also expect small changes in household real incomes. At the national level all income deciles gain except the first, second, and fifth deciles. Regardless, the gains from the changes in the gas industry are quite small. The slightly regressive pattern of changes is confirmed by the rise in the national Gini coefficient (0.01%).

Income decile	NSW	Vic	Qld	SA	WA	Tas	NT	ACT	Aust
Lowest	-0.05	0.06	-0.01	-0.09	0.02	0.00	0.02	-0.07	-0.01
Second	0.00	0.02	-0.01	-0.12	0.09	0.00	0.02	-0.09	0.00
Third	0.00	0.06	0.01	-0.09	0.07	-0.01	0.03	-0.16	0.02
Fourth	0.00	0.03	-0.01	-0.08	0.09	0.00	0.01	-0.19	0.01
Fifth	0.02	0.01	-0.02	-0.09	0.09	-0.01	0.08	-0.17	0.00
Sixth	0.04	0.01	-0.02	-0.13	0.09	-0.02	0.07	-0.14	0.01
Seventh	0.04	0.04	-0.02	-0.12	0.09	-0.02	0.02	-0.16	0.02
Eighth	0.05	0.02	-0.04	-0.12	0.08	-0.03	-0.17	-0.13	0.01
Ninth	0.06	0.02	-0.02	-0.13	0.09	-0.02	0.02	-0.11	0.02
Highest	0.10	0.07	-0.01	-0.08	0.09	-0.02	0.05	-0.16	0.05
All deciles	0.04	0.04	-0.02	-0.11	0.09	-0.02	0.02	-0.14	0.02
Gini coefficient	0.02	0.01	0.00	0.00	0.00	0.00	-0.01	0.00	0.01

Table 6 Changes in household real income and inequality (percentage change)

Source: MMRF-ID simulation.

The first two deciles experience lower or stagnant real incomes because of lower government benefits (due to the reductions in the number of unemployed) and lower labour income for occupations that experience lower wages: Tradespersons, Clerks and LRWs. The fifth decile experiences stagnant real incomes due lower labour income for Clerks and Paraprofessionals.

The national pattern of slightly regressive results is only replicated in NSW and Victoria; the Gini coefficient is constant or falls slightly in all other regions. Although not reported, the detailed results show price effects to be just as important as the income effects in determining the changes in real income for most deciles. Thus, the direct and indirect effects are of about equal size.

### 6. Sensitivity analysis

It is possible that our results are sensitive to many assumptions underlying the implementation of MMRF. Therefore, it is appropriate to investigate the sensitivity of the model results with respect to key assumptions and parameters so as to assess the robustness of the results. We undertake two types of sensitivity analysis: *ad hoc* and systematic.

Ad hoc sensitivity analysis is applied for model assumptions that are binary, e.g., closure choices; Table 7 reports the results of ad hoc sensitivity analysis. These results are relative to our central case results reported in Section 5, so that we vary one set of assumptions from the central case and all leave all other assumptions unchanged. In our central case simulation real government consumption expenditure is a fixed share of real household consumption; real household consumption falls slightly due to the small contractionary effect of the gas industry changes. Holding real expenditure fixed for all governments means more tax revenue must be paid by households via a higher direct tax rate; relative to our central case, this has no impact nationally and slightly benefits households in some regions but not in others. Income inequality is unchanged nationally and in most regions. The ACT experiences a smaller real income loss because of its reliance on federal government spending, which is higher when government expenditure is held fixed. Letting all government budget deficits vary and fixing all tax rates means budget deficits in most regions rise and government debt increases. This removes the stimulatory effect of lower payroll tax rates for most regions observed in the central case. Thus, real output rises by less in most regions and real income is slightly lower nationally and for most regions.

Variable	NSW	Vic	Qld	SA	WA	Tas	NT	ACT	Aust
				1.	Central ca	ise			
All deciles	0.04	0.04	-0.02	-0.11	0.09	-0.02	0.02	-0.14	0.02
Gini coefficient	0.02	0.01	0.00	0.00	0.00	0.00	-0.01	0.00	0.01
			2.	Exogenou	s governm	ent spendi	ng		
All deciles	0.05	0.03	-0.02	-0.14	0.10	-0.03	0.03	-0.12	0.02
Gini coefficient	0.02	0.00	0.00	0.00	0.00	-0.01	-0.01	0.00	0.01
		3. <u>End</u>	ogenous fe	deral & sta	te budget	deficits, ex	kogenous ta	ax rates	
All deciles	0.05	0.02	-0.02	-0.14	0.10	-0.02	0.00	-0.17	0.01
Gini coefficient	0.02	0.00	0.00	-0.01	0.01	0.00	-0.01	0.00	0.01
				4. Verikio	os and Zha	ng (2008)			
All deciles	0.00	0.02	0.00	0.00	0.00	0.00	0.01	0.09	0.01
Gini coefficient	0.03	0.02	0.02	0.08	0.00	0.02	0.02	0.06	0.02

 Table 7 Results of ad hoc sensitivity analysis: household real income and inequality (percentage change)

Source: MMRF-ID simulation.

Table 7 also presents the results of Verikios and Zhang (2008), of which this work is an extension. In many respects our analytical framework is very similar to Verikios and Zhang's (2008). Nevertheless, we do apply different assumptions in key areas when implementing the gas industry changes in MMRF; thus, Verikios and Zhang (2008) presents another example of sensitivity analysis. Relative to our central case, Verikios and Zhang's (2008) results show three main differences: one, the overall increase in real income is smaller; two, the income effects across regions are more evenly distributed; three, income inequality is higher.

In Verikios and Zhang (2008) changes in government revenue from productivity changes are not fully returned to households in the form of changes in direct (payroll) taxes applied by the federal (regional) government(s). A portion of federal government revenue is reduced by lower government benefits due to indexation to wage rates, which themselves fall in response to the productivity changes. This causes lower income decile households to benefit less than higher deciles from the gas industry changes; so income inequality is higher. Verikios and Zhang (2008) also assume labour moves between regions to eliminate regional wage differences, so productivity changes in one region are largely spread evenly across all regions via labour mobility. In contrast, we assume labour market equilibrium operates at the regional-occupational level [see equation (7)], so productivity changes in one regions in our work.

Systematic sensitivity analysis is undertaken to find out how sensitive the results are to variations in key model parameters (elasticities). Table 8 reports the estimated means and standard deviations for real household income and inequality if the relevant parameters vary

symmetrically by up to 50% following a triangular distribution. The calculation of means and standard deviations was carried out using the systematic sensitivity methods automated in the GEMPACK economic modelling software (Harrison and Pearson 1996). These methods rely on a Gaussian quadrature to select a modest number of different sets of values for the varying parameters (DeVuyst and Preckel 1997). The model is solved using each different set of parameter values and the means and standard deviations are calculated over the several different solves of the model. The calculated means and standard deviations will be good approximations to the true means and standard deviations provided that: (i) simulation results are well approximated by a third-order polynomial in the varying shocks and parameters; (ii) varying shocks and parameters have a symmetric distribution; (iii) shocks and parameters do not both vary at once; (iv) shocks and parameters either have a zero correlation or are perfectly correlated within a specified range chosen by the user (e.g.,  $\pm 50\%$ ) (Arndt and Pearson 1996).

In Table 8 the first two rows are the calculated means across the different solves. As expected they are the same as for the original simulation as reported in Table 7. The other sets of results in Table 8 report the values of the standard deviations as each group of parameters (e.g., elasticity of substitution between occupations) is varied by 50%. When calculating means and standard deviations, the industry/commodity dimension of each parameter value is varied together whereas the regional dimension is varied independently.<sup>16</sup> The results indicate that our estimates of household real income effects are remarkably robust with respect to variations in nearly all model parameters because the estimated standard deviations are much smaller than the simulation results. The only exception to this is the elasticity of primary factor substitution. In this case the standard deviations are not insignificant compared to the size of the model results, which means that the results are quite sensitive to the values of this parameter. The results also show our estimates of inequality are invariant to model parameters. Thus, we can be fairly confident of the size of the overall effect on households' welfare and inequality, at the regional and national level, from the estimated changes in the gas industry.

<sup>&</sup>lt;sup>16</sup> For example, in testing the sensitivity with respect to the elasticity of substitution between occupations, regional variations were independent (so the elasticity may be varied up in one region and down in another region in one of the simulations) but are varied together for all industries in each region. We used Stroud's quadrature, which requires running 16 ( $=2\times8$  regions) simulations.

Variable	NSW	Vic	Qld	SA	WA	Tas	NT	ACT	Aust
					1. <u>Mean</u>				
All deciles	0.04	0.04	-0.02	-0.11	0.09	-0.02	0.02	-0.14	0.02
Gini coefficient	0.02	0.01	0.00	0.00	0.00	0.00	-0.01	0.00	0.01
_	2. Elasticity of substitution between occupations								
All deciles	0.000	0.001	0.000	0.000	0.000	0.000	0.000	0.003	0.000
Gini coefficient	0.001	0.002	0.000	0.001	0.000	0.000	0.000	0.001	0.001
_			3. <u>Ela</u>	sticity of p	orimary fac	ctor substit	tution		
All deciles	0.012	0.010	0.014	0.014	0.016	0.010	0.025	0.029	0.012
Gini coefficient	0.003	0.005	0.003	0.005	0.005	0.003	0.003	0.004	0.004
			4. <u>Elas</u>	sticity of in	nport-dom	estic subst	itution		
All deciles	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Gini coefficient	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
			5. <u>Ela</u>	sticity of i	ntra-dome	stic substi	tution		
All deciles	0.001	0.001	0.002	0.001	0.002	0.004	0.005	0.002	0.000
Gini coefficient	0.000	0.000	0.001	0.001	0.001	0.001	0.001	0.000	0.000
				6. Elastici	ty of expo	rt demand			
All deciles	0.001	0.003	0.000	0.000	0.004	0.001	0.004	0.000	0.001
Gini coefficient	0.000	0.001	0.000	0.000	0.001	0.001	0.001	0.000	0.000
		7.	Elasticity of	of employi	nent with 1	respect to	the real wa	ge	
All deciles	0.001	0.002	0.001	0.004	0.002	0.001	0.002	0.009	0.000
Gini coefficient	0.000	0.001	0.001	0.003	0.001	0.001	0.000	0.000	0.001
				8. Elastic	ity of labo	ur supply			
All deciles	0.000	0.000	0.000	0.001	0.000	0.000	0.000	0.001	0.000
Gini coefficient	0.000	0.000	0.000	0.001	0.000	0.000	0.000	0.000	0.000

 Table 8 Results of systematic sensitivity analysis: household real income and inequality (percentage change)

*Source:* MMRF-ID simulation.

The two kinds of sensitivity analysis carried out (*ad hoc* for the model structure, and systematic for the model parameters) suggest that our results are largely a product of our data and the size of the estimated changes in the gas industry rather than the model structure and parameters.

### 7. Concluding remarks

We apply a simple framework for analysing the distributional impacts of structural changes in the national or regional economies of Australia. The framework combines an existing general equilibrium model and a microsimulation model, with detailed household income and expenditure data, to analyse the direct and indirect effects on household income owing to structural change. Applying the technique to the gas industry, our results show that changes in the industry over the 1990s have had minor impacts on households, in terms of real income and inequality. Overall, household real income is higher by 0.02%. This hides the uneven distribution of the effects across regions; households in NSW, Victoria and WA benefit slightly whereas they lose slightly in Queensland, SA, Tasmania and ACT. For most regions inequality is unchanged or slightly higher. Nationally, the Gini coefficient is estimated to have increased slightly by 0.01%.

This work makes a number of contributions. One, it adds to the few Australian studies that have attempted to estimate the distributional effects of structural changes due to microeconomic reform motivated by the Hilmer Report. Two, it represents a methodological advance on these existing studies by estimating the effects on both sides of the household budget, i.e., the expenditure effects and the income effects.

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